

Def. HILBERT SPACE

$H$  is a vectorial space on  $\mathbb{R}$  (or on  $\mathbb{C}$ )

$\|\cdot\|_H$  norm  $d_H(x, y) = \|x - y\|_H$

$H$  is Banach with respect to the convergence associated with  $d_H$

On  $H$  is defined a SCALAR PRODUCT

$$\langle \cdot, \cdot \rangle : H \times H \longrightarrow \mathbb{R}$$

1)  $\langle x, y \rangle = \langle y, x \rangle$  SYMMETRIC

2) LINEAR  $\langle \lambda x + \mu z, y \rangle =$   
 $= \lambda \langle x, y \rangle + \mu \langle z, y \rangle$

$$\begin{aligned} \lambda, \mu &\in \mathbb{R} \\ x, z, y &\in H \end{aligned}$$

### 3) CONTINUOUS

if  $x_n \rightarrow x$  in  $H$

$$(d_H(x_n, x) = \|x_n - x\|_H \rightarrow 0 \text{ as } n \rightarrow \infty)$$

$\forall y \in H$

$$\langle x_n, y \rangle \xrightarrow{n \rightarrow \infty} \langle x, y \rangle$$

### 4) CAUCHY - SCHWARTZ

$$|\langle x, y \rangle| \leq \|x\|_H \|y\|_H$$

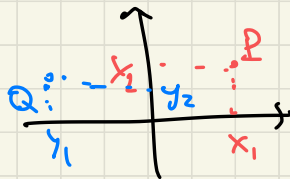
5) it is associated with the norm

$$\langle x, x \rangle = \|x\|_H^2$$

~ . ~

$$\mathcal{E}_X \quad \mathbb{R}^2 = \{ (x_1, x_2) \mid x_i \in \mathbb{R} \}$$

$$\langle \underbrace{(x_1, x_2)}_P, \underbrace{(y_1, y_2)}_Q \rangle = x_1 y_1 + x_2 y_2$$



$$\mathbb{R}^n \quad \langle (x_1 \dots x_n), (y_1 \dots y_n) \rangle = x_1 y_1 + x_2 y_2 + \dots + x_n y_n$$

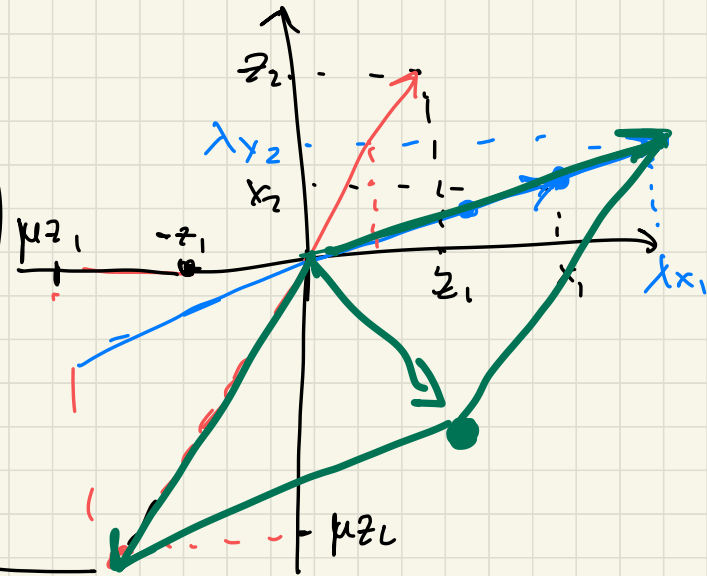
$$\underline{\lambda}(x_1, x_2) + \underline{\mu}(z_1, z_2) = (\lambda x_1 + \mu z_1, \lambda x_2 + \mu z_2) \quad \mu < 0$$

$$\langle \underline{\lambda}(x_1, x_2) + \underline{\mu}(z_1, z_2), \underline{(y_1, y_2)} \rangle =$$

$$= \lambda x_1 y_1 + \mu z_1 y_1 + \lambda x_2 y_2 + \mu z_2 y_2$$

$$|\langle (x_1, x_2), (y_1, y_2) \rangle| =$$

$$= |x_1 y_1 + x_2 y_2| \leq \underbrace{\sqrt{x_1^2 + x_2^2}}_{|(x_1, x_2)|} \underbrace{\sqrt{y_1^2 + y_2^2}}_{|(y_1, y_2)|}$$



$$(x_1^n, x_2^n) \rightarrow (x_1, x_2)$$

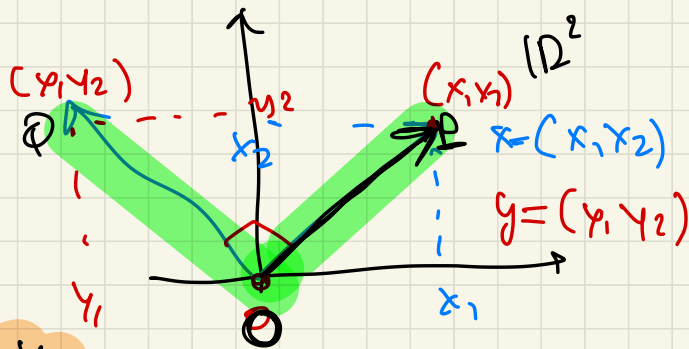
$$x_1^n \rightarrow x_1$$

$$x_2^n \rightarrow x_2$$

The notion of scalar product permits to introduce the DEFINITION of ORTHOGONALITY

Def:  $x \perp y$  ( $x, y \in H$ )  $x$  is orthogonal to  $y$

$$\text{iff } \langle x, y \rangle = 0$$



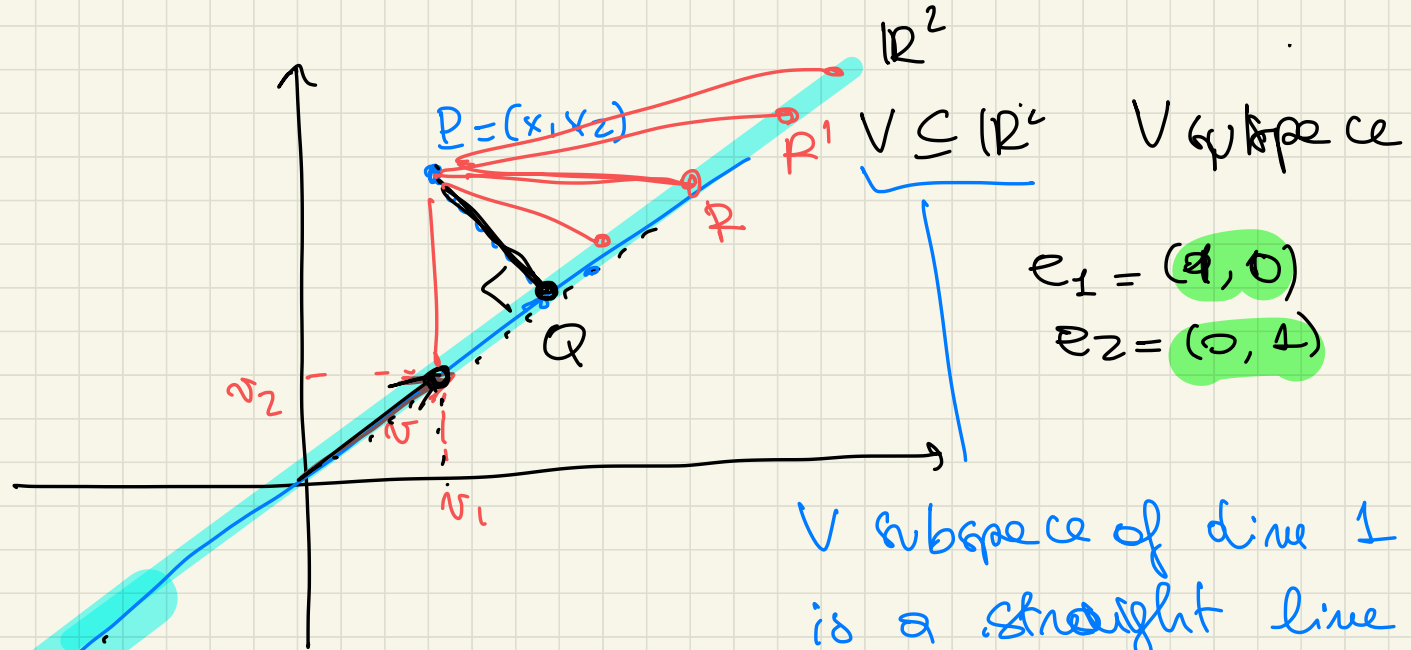
$$\langle (x_1, x_2), (y_1, y_2) \rangle = 0 = x_1 y_1 + x_2 y_2$$

$$\Leftrightarrow \overline{OP} \perp \overline{OQ}$$

$$P = (x_1, x_2) \quad Q = (y_1, y_2)$$

$$O = (0, 0)$$

The main theorem in Hilbert spaces is the  
ORTHOGONAL PROJECTION THEOREM



$V$  subspace of line 1  
 is a straight line  
 passing through  $O$ .

$Q: \exists Q \in V$  such that  
 $|Q - P| = \min_{R \in V} |R - P|$

$$V = \{ \lambda (v_1, v_2), \lambda \in \mathbb{R} \}$$

# ORTHOGONAL PROJECTION THEOREM

$H$  Hilbert

$V \subseteq H$  SUBSPACE, CLOSED

( $V$  is a vectorial space CONTAINED in  $H$ )

$$v_1, v_2 \in V \quad \lambda v_1 + \mu v_2 \in V \quad \forall \lambda, \mu \in \mathbb{R}$$

CLOSED  $\left[ \begin{array}{l} v_n \in V \quad v_n \rightarrow x \text{ in } H \quad v_n \text{ is a converging} \\ \text{sequence in } H \Rightarrow x \in V \end{array} \right.$

( $V$  contains all the limits of its converging sequences)

then the following holds:

1)  $\forall h \in H$  there exists a UNIQUE  $v \in V$

such that

$$\|v - h\| = \min_{w \in V} \|w - h\|$$

(There exists a UNIQUE  $v \in V$  which has MINIMAL DISTANCE from  $h$ , AMONG ALL ELEMENTS IN  $V$ )

$$(2) \quad \underline{h-v \in V^\perp} \quad \left( \begin{array}{l} \text{so } \langle h-v, w \rangle = 0 \\ \langle v-h, w \rangle = 0 \\ \forall w \in V \end{array} \right)$$

Therefore  $h = \underbrace{(h-v)}_{\in V^\perp} + \underbrace{v}_{\in V}$

↓  
Every element of  $H$  can be written in a UNIQUE WAY as the sum of one element of  $V$  and one element of  $V^\perp$ .

No proof.

Just prove of the fact that (1)  $\Rightarrow$  (2)

if  $v \in V$  is the element at minimal distance  
from  $h$   $\|h-v\| = \min_{w \in V} \|h-w\|$

$$\Rightarrow h-v \perp V \quad (h-v) \in V^\perp$$

$$x \in \mathbb{R} \quad w \in V \quad \underline{v + xw} \in V$$

$$\|h-v\|^2 \leq \|h-(v+xw)\|^2 \quad \forall x$$

$$\underbrace{\langle h-v, h-v \rangle} \leq \underbrace{\langle h-(v+xw), h-(v+xw) \rangle} \quad \forall x \in \mathbb{R}$$



$$\phi(z) = \langle h - (\sigma + \pi w), h - (\sigma + \pi w) \rangle$$

$$\phi(0) \leq \phi(z) \quad \forall z$$

↓

$$\phi'(0) = 0$$

(since  $\pi=0$  is a minimum point)

↓

$$\phi(z) = \langle h - \sigma - \pi w, h - \sigma - \pi w \rangle =$$

$$= \langle h - \sigma, h - \sigma \rangle + \langle -\pi w, h - \sigma \rangle +$$

$$+ \langle h - \sigma, -\pi w \rangle + \langle -\pi w, -\pi w \rangle =$$

$$= \|h - \sigma\|^2 - 2\pi \langle h - \sigma, w \rangle + \pi^2 \|w\|^2$$

$$\phi'(z) = -2 \langle h - \sigma, w \rangle + 2\pi \|w\|^2$$

$$\phi'(0) = -2 \langle h - \sigma, w \rangle = 0$$

$$w \in V \Rightarrow h - \sigma \perp V$$

Let us fix a PROBABILITY SPACE

$\Omega$  set

$\mathcal{F}$   $\sigma$ -algebra on  $\Omega$  (filtration)

$\mathbb{P}$  probability measure

$L^2 = \{ X : \Omega \rightarrow \mathbb{R} \text{ random variable such that } \underbrace{\mathbb{E}(|X|^2)}_{//} < +\infty \}$

$$\mathbb{E}(|X|^2) = \int_{\mathbb{R}} x^2 d\mathcal{L}_X(x)$$

$$\mathcal{L}_X = X \# \underline{\mathbb{P}}$$

$$\mathcal{L}_X(A) = \mathbb{P}(\{\omega \mid X(\omega) \in A\})$$

$A \in \mathcal{B}(\mathbb{R})$

$$\text{On } M^2 \quad \|X\|_2 = \text{norm} = \sqrt{\mathbb{E}(|X|^2)} = \sqrt{\int_{\mathbb{R}} z^2 d\mathcal{L}_X(z)}$$

$\downarrow$   
 distance  $X, Y$

$$\|X - Y\|_2 = \sqrt{\mathbb{E}(|X - Y|^2)} = \sqrt{\int_{\mathbb{R} \times \mathbb{R}} |x - y|^2 d\mathcal{L}_{(X, Y)}(x, y)}$$

$$X, Y \in M^2$$

$$X: \Omega \rightarrow \mathbb{R}$$

$$Y: \Omega \rightarrow \mathbb{R}$$

$$(X, Y): \Omega \rightarrow \mathbb{R} \times \mathbb{R} = \mathbb{R}^2$$

$$\omega \mapsto (X(\omega), Y(\omega))$$

$\downarrow$  I define the joint law

$\mathcal{L}_{(X, Y)}$  a measure in  $\mathbb{R}^2$

$$\mathcal{L}_{(X, Y)} = (X, Y) \# \mathbb{P}, \quad A, B \in \mathcal{B}(\mathbb{R}) \quad A \times B \in \mathcal{B}(\mathbb{R}^2)$$

$$\mathcal{L}_{(X, Y)}(\underline{A} \times \underline{B}) = \mathbb{P} \{ \omega \in \Omega, \underline{X}(\omega) \in A, \underline{Y}(\omega) \in B \}$$

if  $X$  and  $Y$  are independent

$$L_{(X,Y)}(x,y) = L_X(x) L_Y(y) .$$

~ . ~

In  $M^2$  I define the scalar product

$$\langle X, Y \rangle = \mathbb{E}(XY) = \int_{\mathbb{R} \times \mathbb{R}} xy \, dL_{(X,Y)}(x,y)$$

Note that it satisfies all the properties listed at the beginning

$$\mathbb{E}(X X) = \mathbb{E}(X)^2 = \|X\|^2$$

Recall  $X \perp Y \Leftrightarrow \mathbb{E}(XY) = 0$

$V \subseteq M^2 \Rightarrow V^\perp = \{ Y \in M^2 \text{ such that } \mathbb{E}(XY) = 0 \}$   
 $\forall X \in V$

Ex  $V = \{ Y \in M^2 \mid \mathbb{E}(Y) = 0 \}$   $V^\perp = \{ \text{constant random variables} \}$

Fix  $X \in M^2$ .

What is the orthogonal projection of  $X$  in  $V$ ?

What is the orthogonal projection of  $X$  in  $V^\perp$ ?

$$X = X - \mathbb{E}(X) + \mathbb{E}(X)$$

Note that  $\mathbb{E}(X)$  is a constant  $\Rightarrow \mathbb{E}(X) \in V^\perp$

$X - \mathbb{E}(X) \in V$  since  $\mathbb{E}(X - \mathbb{E}(X)) = \mathbb{E}(X) - \mathbb{E}(X) = 0$

By the orthogonal projection theorem

$X - \mathbb{E}(X)$  is the orthogonal projection of  $X$  in  $V$   
(it is the random variable with 0 mean  
that is at minimal distance from  $X$ ,  
that is: it is the best possible approximation of  $X$   
by a random variable with 0 mean

$\mathbb{E}(X)$  is the orthogonal projection of  $X$  in  $V^\perp$

$\mathbb{E}(X)$  is the constant which is at minimal distance from  $X$  among all other constants

$$\min_{c \in \mathbb{R}} \mathbb{E}(X - c)^2 = \mathbb{E}(X - \mathbb{E}(X))^2.$$

Es  $M^2 = \{X : (\Omega, \mathcal{F}, \mathbb{P}) \rightarrow \mathbb{R} \text{ random variable} \\ E(X^2) < +\infty\}$

Let  $\mathcal{G} \subsetneq \mathcal{F}$  a  $\sigma$ -algebra contained in  $\mathcal{F}$

$M^2_{\mathcal{G}} = \{X : (\Omega, \mathcal{G}, \mathbb{P}) \rightarrow \mathbb{R}, \text{ random variable} \\ E(X^2) < +\infty\}$

$M^2_{\mathcal{G}}$  are random variables  $\mathcal{G}$ -measurable

$X \in M^2_{\mathcal{G}} \Leftrightarrow \forall x \in \mathbb{R} \quad \{\omega \in \Omega : X(\omega) \leq x\} \in \mathcal{G} \subseteq \mathcal{F} \Rightarrow \\ X \in M^2 \quad (M^2_{\mathcal{G}} \subseteq M^2)$

$M^2_{\mathcal{G}} \subsetneq M^2$  : let  $A \in \mathcal{F} \setminus \mathcal{G}$  (then also  $\Omega \setminus A \in \mathcal{F} \setminus \mathcal{G}$ )

$1_A : \Omega \rightarrow \mathbb{R} \quad 1_A(\omega) = \begin{cases} 0 & \omega \notin A \\ 1 & \omega \in A \end{cases}$  then  $1_A \in M^2 \quad 1_A \notin M^2_{\mathcal{G}}$

$$X \in M^2$$

$$X: (\Omega, \mathcal{F}, \mathbb{P}) \rightarrow \mathbb{R}$$

↓ orthogonal projection of  $X$  in  $M^2_{\mathcal{G}}$  is a

RANDOM VARIABLE in  $M^2_{\mathcal{G}}$  (measurable w.r.t.  $\mathcal{G}$ )

which is **the best approximation** of  $X$   
among  $\mathcal{G}$ -measurable random variables

↓  $E(X | \mathcal{G})$  conditional expectation of  $X$   
on  $\mathcal{G}$ .

(the "least square" estimator of  $X$  among  
 $\mathcal{G}$ -measurable random variables).



$$\mathbb{E}(X|Y) \in M^2_Y \text{ and}$$

$$\mathbb{E}(|X - \mathbb{E}(X|Y)|^2) = \min_{Y \in M^2_Y} \mathbb{E}(|X - Y|^2)$$

$$\text{So } \nexists X \in M^2_Y \Rightarrow \mathbb{E}(X|Y) = X$$

$X - \mathbb{E}(X|Y)$  is orthogonal to  $M^2_Y$

$$\Rightarrow \forall Z \in M^2_Y \quad \mathbb{E}((X - \mathbb{E}(X|Y))Z) = 0 = \mathbb{E}(XZ) - \mathbb{E}(\mathbb{E}(X|Y)Z)$$

$$\Rightarrow \mathbb{E}(XZ) = \mathbb{E}(\mathbb{E}(X|Y)Z) \quad \forall Z \in M^2_Y$$

$$Z = c \text{ constant } \in M^2_Y \quad \text{since } \{\omega \mid \mathbb{P}(\omega) \leq x\} = \begin{cases} \Omega & x \geq c \\ \emptyset & x < c \end{cases}$$

take  $c = 1$

$$\mathbb{E}(X) = \mathbb{E}(\mathbb{E}(X|Y))$$

Ex Let us fix  $Y \in M^2$ .

$\sigma(Y)$  =  $\sigma$ -algebra contained in  $\mathcal{F}$  generated  
by  $\{\omega, Y(\omega) \leq x\} \quad \forall x \in \mathbb{R}$   
= minimal  $\sigma$ -algebra which makes  $Y$  measur.

$M^2_{\sigma(Y)} = \{ \text{random variables in } M^2 \text{ which are} \\ \text{measurable with respect to } \sigma(Y) \}$   
(IT IS POSSIBLE TO PROVE)

$= \{ R(Y) \quad h: \mathbb{R} \rightarrow \mathbb{R} \text{ measurable} \\ \text{such that } h(Y): \Omega \longrightarrow \mathbb{R}$

is a random variable  
and  $E(R(Y))^2 < +\infty \}$

$\omega \xrightarrow{Y} Y(\omega) \xrightarrow{h} h(Y(\omega))$   
 $\Omega \xrightarrow{Y} \mathbb{R} \xrightarrow{h} \mathbb{R}$

Example take  $Y$ : tossing a coin

$$Y(\omega) = \begin{cases} 1 & \text{if } \omega \text{ is head} \\ 0 & \text{if } \omega \text{ is tail} \end{cases}$$

$$\sigma(Y) = \{ \Omega, \emptyset, \underbrace{\{\omega \mid Y(\omega) = 1\}}_{\sim}, \underbrace{\{\omega \mid Y(\omega) = 0\}}_{\sim} \}$$

For  $Y \in M^2$  we define

$\mathbb{E}(X|Y)$  = conditional expectation of  $X$   
given  $Y =$

$= \mathbb{E}(X|\sigma(Y))$  orthogonal projection of  
 $X$  in  $M^2_{\sigma(Y)}$

$\mathbb{E}(X|Y)$  is the random variable  $h(Y)$  which  
best approximate  $X$ .

$$\mathbb{E}(X|Y) \in M^2_{\sigma(Y)} \quad \text{so} \quad \mathbb{E}(X|Y) = h(Y)$$

$$h(y) = \mathbb{E}(X | Y=y)$$

↓      ...

how to compute  $\mathbb{E}(X|Y)$ ?

$$\min_{\substack{z \in M^2_{\sigma(Y)}}} \mathbb{E}(X-z)^2 = \min_{\substack{f: \mathbb{R} \rightarrow \mathbb{R}^2 \\ \text{meas.}}} \mathbb{E}(X-f(Y))^2$$

$\Rightarrow$  since  $z \in M^2_{\sigma(Y)} \Leftrightarrow z = f(Y) \quad f: \mathbb{R} \rightarrow \mathbb{R}$

$h: \mathbb{R} \rightarrow \mathbb{R}$  such that

$$\mathbb{E}(X-h(Y))^2 = \min_{f: \mathbb{R} \rightarrow \mathbb{R}} \mathbb{E}(X-f(Y))^2$$

then  $h(Y) = \mathbb{E}(X|Y)$       LEAST SQUARE ESTIMATOR of  $X$  given  $Y$ .

solving the minimization problem is quite difficult!

In some case it is easy:

take  $X$  independent of  $Y \Rightarrow$  so  $E(X \cdot f(Y)) = E(X) E(f(Y))$   
 $\forall f: \mathbb{R} \rightarrow \mathbb{R}$

Observe that  $X - E(X) \perp M_{\sigma(Y)}^2$  since

$$\forall z \in M_{\sigma(Y)}^2 \quad z = f(Y) \quad f: \mathbb{R} \rightarrow \mathbb{R}$$

$$E((X - E(X)) f(Y)) = E(X f(Y) - E(X) f(Y)) = \text{(by independ.)}$$

$$= E(X) E(f(Y)) - E(X) E(f(Y)) = 0$$

Moreover  $E(X) \in M_{\sigma(Y)}^2$  since it is a constant

$$X = \underbrace{E(X)}_{M_{\sigma(Y)}^2} + \underbrace{X - E(X)}_{(M_{\sigma(Y)}^2)^\perp} \Rightarrow E(X|Y) = E(X) \text{ constant!}$$

In general though computing  $\mathbb{E}(X|Y)$  is difficult

$$\min_{f: \mathbb{R} \rightarrow \mathbb{R}} \mathbb{E}(X - f(Y))^2 = \mathbb{E}(X - \mathbb{E}(X|Y))^2$$

$\downarrow$   
 $\mathbb{E}(X - a(Y))^2 \quad a(Y) = \mathbb{E}(X|Y)$

I restrict the set where I compute the minimum

I consider just  $f: \mathbb{R} \rightarrow \mathbb{R} \quad f(x) = ax + b$   
 $f$  LINEAR.

$V = \{ \underbrace{ax + b}_{\text{the orthogonal projection of } x \text{ in } V} \mid a, b \in \mathbb{R} \} \subsetneq M^2_{\mathbb{R}}(Y)$

is not  $\mathbb{E}(X|Y)$  but it is just the

best LINEAR APPROXIMATION of  $X$  given  $Y$   
(the LINEAR FUNCTION of  $Y$  which is at MINIMAL  
distance from  $X$ )  
 $\bar{a}, \bar{b} \in \mathbb{R}$

$$\mathbb{E}|X - \bar{a}Y - \bar{b}|^2 = \min_{(a,b) \in \mathbb{R}^2} \mathbb{E}(\underbrace{|X - aY - b|}^2)$$

Ex: let us find  $\bar{a}, \bar{b}$

$$\min_{a,b} \mathbb{E}((X - aY - b)^2) = \min_{a,b} (\mathbb{E}(X^2) + a^2 \mathbb{E}(Y)^2 + b^2 +$$

$$- 2a \mathbb{E}(XY) - 2b \mathbb{E}(X) + 2ab \mathbb{E}(Y)).$$

$$\phi(a, b) := \mathbb{E}(x^2) + a^2 \mathbb{E}(y^2) + b^2 - 2a \mathbb{E}(xy) - 2b \mathbb{E}(x) + 2ab \mathbb{E}(y)$$

↓  
compute  $\min_{a, b} \phi(a, b)$

$$\begin{cases} \frac{\partial \phi}{\partial a}(a, b) = 2a \mathbb{E}(y^2) - 2 \mathbb{E}(xy) + 2b \mathbb{E}(y) = 0 \\ \frac{\partial \phi}{\partial b}(a, b) = 2b - 2 \mathbb{E}(x) + 2a \mathbb{E}(y) = 0 \end{cases}$$

$$\begin{cases} a = \frac{\mathbb{E}(xy) - b \mathbb{E}(y)}{\mathbb{E}(y^2)} = \frac{\mathbb{E}(xy)}{\mathbb{E}(y^2)} - \frac{\mathbb{E}(x) \mathbb{E}(y)}{\mathbb{E}(y^2)} + a \frac{(\mathbb{E}(y))^2}{\mathbb{E}(y^2)} \\ b = \mathbb{E}(x) - a \mathbb{E}(y) \end{cases}$$



$$\Rightarrow \begin{cases} a \left( 1 - \frac{(\mathbb{E}(Y))^2}{\mathbb{E}(Y^2)} \right) = \frac{\text{Cov}(X, Y)}{\mathbb{E}(Y^2)} \\ b = \mathbb{E}(X) - a \mathbb{E}(Y) \end{cases}$$

$$\text{Cov}(X, Y) = \mathbb{E}(XY) - \mathbb{E}(X)\mathbb{E}(Y)$$

$$\text{Var}(Y) = \mathbb{E}(Y^2) - (\mathbb{E}(Y))^2$$

$$\Rightarrow a \left( \frac{\mathbb{E}(Y^2) - (\mathbb{E}(Y))^2}{\cancel{\mathbb{E}(Y^2)}} \right) = \frac{\cancel{\text{Cov}(X, Y)}}{\cancel{\mathbb{E}(Y^2)}} \rightarrow a = \frac{\text{Cov}(X, Y)}{\text{Var } Y}$$

$$b = \mathbb{E}(X) - \frac{\text{Cov}(X, Y)}{\text{Var } Y} \cdot \mathbb{E}(Y)$$

$L(X|Y)$  = the linear least square estimator of  $X$  given  $Y$

$$= \frac{\text{Cov}(X, Y)}{\text{Var } Y} Y + \mathbb{E}(X) - \frac{\text{Cov}(X, Y)}{\text{Var } Y} \mathbb{E}(Y).$$

$$a = \frac{\cancel{E(Y^2)} [E(XY) - E(X)E(Y)]}{\cancel{E(Y^2)} [E(Y^2) - \{E(Y)\}^2]} = \frac{\text{Cov}(X, Y)}{\text{Var}(Y)}$$

$$b = E(X)E(Y^2) - E(XY)E(Y)$$

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y)$$

$$\text{Var } Y = E(Y^2) - E(Y)^2$$