

# CONDITIONAL EXPECTATION of $X$ with respect to a $\sigma$ -algebra

$(\Omega, \mathcal{F}, \mathbb{P})$  probability space

$$M^2 = \{ X : (\Omega, \mathcal{F}, \mathbb{P}) \rightarrow \mathbb{R} \text{ measurable} \\ E(X^2) < +\infty \}$$

$\mathcal{G} \subseteq \mathcal{F}$   $\mathcal{G}$   $\sigma$ -algebra contained in  $\mathcal{F}$

$$M^2_{\mathcal{G}} = \{ Y \in M^2 \text{ such that } \forall B \in \mathcal{B}(\mathbb{R}) \quad Y^{-1}(B) \in \mathcal{G} \}$$

$E(X | \mathcal{G}) =$  Conditional expectation of  $X$  with respect to  $\mathcal{G} =$  orthogonal projection of  $X$  in  $M^2_{\mathcal{G}}$

1)  $E(X | \mathcal{G}) \in M^2_{\mathcal{G}}$       2)  $E[(X - E(X | \mathcal{G}))^2] = \min_{Y \in M^2_{\mathcal{G}}} E[(X - Y)^2]$

3)  $\forall Y \in M^2_{\mathcal{G}} \quad X - E(X | \mathcal{G}) \perp Y$   
which means  $E((X - E(X | \mathcal{G}))Y) = 0$

# Conditional expectation with respect to a random variable

$E(X|Y) = E(X | \sigma(Y))$  where  $\sigma(Y)$  is the smallest

$\sigma$ -algebra which contains all elements  $Y^{-1}(B)$ ,  
for  $B \in \mathcal{B}(\mathbb{R})$ .

$M^2_{\sigma(Y)} = \{ g(Y) \mid \text{for some } g: \mathbb{R} \rightarrow \mathbb{R} \text{ such that } g(Y) \in M^2 \}$   
so  $g$  measurable and  $E(g(Y))^2 < +\infty$

$E(X|Y)$  is the orthogonal projection on  $M^2_{\sigma(Y)}$

1)  $E(X|Y) = h(Y) \quad \exists h: \mathbb{R} \rightarrow \mathbb{R}$

2)  $\min_{\substack{g: \mathbb{R} \rightarrow \mathbb{R} \\ \text{meas.}}} E(X - g(Y))^2 = E(X - E(X|Y))^2$

3)  $E((X - E(X|Y))g(Y)) = 0 \quad \forall g: \mathbb{R} \rightarrow \mathbb{R} \text{ measurable}$

$g$  is particular if  $g(r) = c$

$$E((X - E(X|Y))c) = 0 \Rightarrow E(X) = E(E(X|Y))$$

if  $g(r) = r$

$$E((X - E(X|Y))Y) = 0 \Rightarrow E(XY) = E(E(X|Y) \cdot Y)$$

so if  $X$  and  $Y$  are independent  $E(X|Y) = \underline{E(X)}$  constant

$E(X|Y)$  is the best predictor of  $X$  given  $Y$ .

↓  
since it is difficult in general to find  $E(X|Y) (= g(Y))$

we consider an easier problem:

let  $V = \{aY + b, a, b \in \mathbb{R}\} \subseteq M^2_{\sigma(Y)}$

$V$  is a finite dimensional subspace of  $M^2_{\sigma(Y)}$ , given by

the linear functions of  $Y$ .

$$V \xrightarrow{\sim} \mathbb{R}^2$$
$$aY+b \longmapsto (\bar{a}, \bar{b})$$

$$\mathbb{R}^2 \longrightarrow V$$
$$(\bar{a}, \bar{b}) \longmapsto aY+b$$

$V$  is 2-dimensional since it is isomorphic as a vectorial space to  $\mathbb{R}^2$ .

$\bar{a}Y + \bar{b}$  is the best LINEAR MEAN SQUARE ESTIMATOR of  $X$  given  $Y$  if

$\bar{a}Y + \bar{b}$  is the orthogonal projection of  $X$  on  $V$ .

1)  $\bar{a}Y + \bar{b} \in V$

2)  $E(X - \bar{a}Y - \bar{b})^2 = \min_{(a, b) \in \mathbb{R}^2} E(X - aY - b)^2$

3)  $E[(X - \bar{a}Y - \bar{b})(aY + b)] = 0 \quad \forall (a, b) \in \mathbb{R}^2$

How to compute the orthogonal projection on  $V$ ?

FIRST METHOD

1) we minimize  $\min_{(a,b) \in \mathbb{R}^2} E(X - aY - b)^2$  and find  $\bar{a}, \bar{b}$  MINIMA

SECOND METHOD

2) general method based on orthonormal basis

Def:  $(e_i)_{i \in I}$  is a ORTHONORMAL BASIS of a Hilbert space if it is a basis (of the vectorial space) and  $\|e_i\| = 1 \forall i$   $(e_i, e_j) = 0 \forall i \neq j$

Let  $H$  be a Hilbert space

and  $V$  be a subspace of  $H$  of FINITE DIMENSION

$\hookrightarrow V$  has a finite basis  $\langle v_1, \dots, v_n \rangle$ .

( $\forall v \in V \quad v = \sum_{i=1}^n a_i v_i$  every element of  $V$  is written as a linear combination of elements of the basis).

Starting from  $\langle v_1, \dots, v_n \rangle$  we may construct an ORTHONORMAL BASIS of  $V$  (by the GRAM-SCHMIDT ORTHONORMALIZATION PROCEDURE)

$$v_1 \rightarrow e_1 = \frac{v_1}{\|v_1\|} \quad \text{so } \|e_1\| = 1 = (e_1, e_1)$$

$$e_2 := \frac{v_2 - (v_2, e_1) e_1}{\|v_2 - (v_2, e_1) e_1\|}$$

Note that  $\|e_2\| = 1 = (e_2, e_2)$

$$(e_2, e_1) = \frac{(v_2, e_1) - (v_2, e_1)(e_1, e_1)}{\|v_2 - (v_2, e_1) e_1\|} = 0$$

$$e_3 := \frac{v_3 - (v_3, e_2) e_2 - (v_3, e_1) e_1}{\|v_3 - (v_3, e_2) e_2 - (v_3, e_1) e_1\|}$$

$$\|e_3\| = 1$$

$$(e_3, e_2) = 0 = (e_3, e_1)$$

$$e_4 := \dots$$

$$\langle v_1, \dots, v_n \rangle \rightarrow \langle e_1, \dots, e_n \rangle$$

where  $\|e_i\| = 1 \quad \forall i$

$$(e_i, e_j) = 0 \quad \forall i \neq j$$

$\forall a \in H$  the orthogonal projection of  $a$  in  $V$  is given by  $\sum_{i=1}^n (a, e_i) e_i$  where  $e_i$  is an orthonormal basis of  $V$ .

So let  $V = \{aY + b, a, b \in \mathbb{R}\} \subseteq M_{\mathbb{R}}^2 \subseteq M^2$

a basis of  $V$  is  $\langle 1, Y \rangle$  (every element of  $V$  is written as  $b \cdot 1 + a \cdot Y$ )

$$v_1 = 1 \quad v_2 = Y$$

$$e_1 = \frac{1}{\|1\|} = \frac{1}{(E(1^2))^{1/2}} = 1$$

$$e_2 = \frac{v_2 - (v_2, e_1) e_1}{\|v_2 - (v_2, e_1) e_1\|}$$

$$e_2 := \frac{Y - E(Y \cdot 1) \cdot 1}{(E((Y - E(Y))^2))^{1/2}}$$

the orthonormal basis of  $V$  is

$$\left\langle 1, \frac{Y - E(Y)}{\sqrt{\text{Var}(Y)}} \right\rangle$$

$$\begin{aligned} \text{Var}(Y) &= E((Y - E(Y))^2) = \\ &= E(Y^2) - (E(Y))^2. \end{aligned}$$

↓

the orthogonal projection of  $X$  in  $V$  is given by

$$(X, e_1) e_1 + (X, e_2) e_2 =$$

$$= E(X \cdot 1) \cdot 1 + E\left(X \cdot \frac{Y - E(Y)}{\sqrt{\text{Var}(Y)}}\right) \frac{Y - E(Y)}{\sqrt{\text{Var}(Y)}} =$$

$$= E(X) + \left[ \frac{E(XY) - E(X)E(Y)}{\text{Var}(Y)} \right] (Y - E(Y)) =$$

$$= \left[ E(X) - \frac{\text{Cov}(X, Y)}{\text{Var}(Y)} E(Y) \right] + \left[ \frac{\text{Cov}(X, Y)}{\text{Var}(Y)} \right] \cdot Y$$

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y)$$



# PRODUCT OF CONVOLUTION

$$L^1(\mathbb{R}) = \left\{ f : \mathbb{R} \rightarrow \mathbb{R} \quad \int_{\mathbb{R}} |f(x)| dx < +\infty \right\} \quad \text{not Hilbert (just Banach)}$$

$$L^2(\mathbb{R}) = \left\{ f : \mathbb{R} \rightarrow \mathbb{R} \quad \int_{\mathbb{R}} |f(x)|^2 dx < +\infty \right\} \quad \text{Hilbert}$$

By Hölder inequality we have that  $f, g \in L^2(\mathbb{R})$

$$f \cdot g \in L^1(\mathbb{R}) \quad \int_{\mathbb{R}} |f(x)g(x)| dx \leq \left( \int_{\mathbb{R}} |f(x)|^2 dx \right)^{1/2} \left( \int_{\mathbb{R}} |g(x)|^2 dx \right)^{1/2}$$

We define another notion of product of functions

PRODUCT of CONVOLUTION:

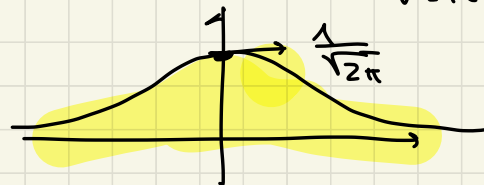
$$(f * g)(x) = \int_{\mathbb{R}} f(x-y)g(y) dy = \int_{\mathbb{R}} f(y)g(x-y) dy$$

~~$f(x)g(x)$~~

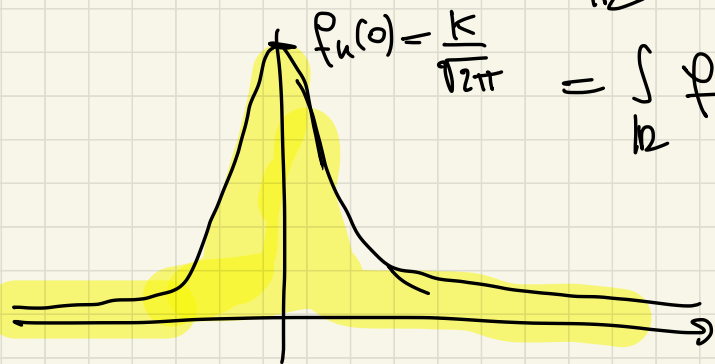
TO UNDERSTAND HOW CONVOLUTION WORKS:

let  $f \geq 0$ ,  $f$  smooth,  $\int_{\mathbb{R}} f(x) dx = 1$  ( $f(x) = \frac{e^{-x^2/2}}{\sqrt{2\pi}}$ )

for  $k \in \mathbb{N}$   $f_k(x) := k f(kx)$   
( $= \frac{k}{\sqrt{2\pi}} e^{-kx^2/2}$ )



then  $f_k \geq 0$   $\int_{\mathbb{R}} f_k(x) dx = \int_{\mathbb{R}} k f(kx) dx = \int_{\mathbb{R}} f(y) dy = 1$  ( $y = kx$   
 $dy = k dx$ )



$$g \in L^1(\mathbb{R}) \quad g * f_k(x) = \int_{\mathbb{R}} g(x-y) f_k(y) dy =$$

$$= \int_{\mathbb{R}} g(x-y) k f(ky) dy = \text{change variable} =$$

$\omega \quad z = ky$   
 $dz = k dy$

$$= \int_{\mathbb{R}} g\left(x - \frac{z}{k}\right) f(z) dz \xrightarrow{k \rightarrow +\infty} g(x) \int_{\mathbb{R}} f(z) dz = g(x)$$

$$g * f_k(x) \xrightarrow{k \rightarrow +\infty} g(x)$$

$g * f_k$  is an "approximation" of  $g$  by taking at every  $x$  the average of the values of  $g$  around  $x$ .

$g * f_k$  is as regular as  $f_k$

since 
$$\frac{d}{dx} (g * f_k)(x) = \int_{\mathbb{R}} g(y) \frac{d}{dx} f_k(x-y) dy$$

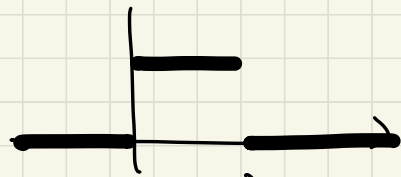
In particular if  $f$  is differentiable  $n$ -times  
 $\Rightarrow g * f$  is differentiable  $n$  times.

Proposition: If  $f, g \in L^2(\mathbb{R})$  then  $f * g \in C(\mathbb{R})$   
and  $\lim_{|x| \rightarrow \infty} f * g(x) = 0$

(NO PROOF / convolution of 2  $L^2$  functions is  
CONTINUOUS and bounded - even if the starting  
functions were not continuous).

Ex example

$$f(x) = g(x) = \begin{cases} 0 & x \leq 0 \\ 1 & 0 < x < 1 \\ 0 & x \geq 1 \end{cases}$$



( $f$  is the density of a uniform random variable).

then  $f * f$  is CONTINUOUS and  $\lim_{|x| \rightarrow +\infty} f * f(x) = 0$ .

(even if  $f$  is NOT CONTINUOUS).

proof

$$f * f(x) = \int_{\mathbb{R}} f(x-y) f(y) dy = \int_{-\infty}^0 f(x-y) f(y) dy +$$

$$+ \int_0^1 f(x-y) f(y) dy + \int_1^{+\infty} f(x-y) f(y) dy =$$

$$= \int_0^1 f(x-y) dy \stackrel{\left( \begin{array}{l} t = x-y \\ dt = -dy \end{array} \right)}{=} \int_x^{x-1} f(t) (-1) dt =$$

$$= \int_{x-1}^x f(t) dt$$

$$f * f(x) = \int_{x-1}^x f(t) dt.$$

1) if  $x < 0$   $(x-1, x) \subseteq (-\infty, 0) \Rightarrow \int_{x-1}^x f(t) dt = 0$

2) if  $0 < x < 1$   $x-1 < 0$  and  $x > 0$

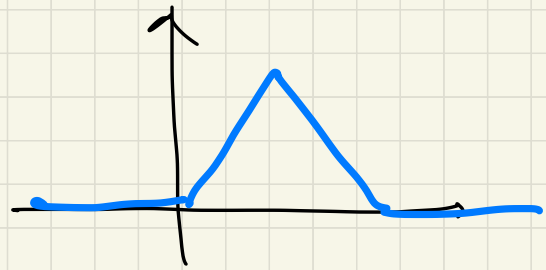
$$\int_{x-1}^x f(t) dt = \int_{x-1}^0 \cancel{f(t) dt} + \int_0^x \cancel{f(t) dt} = \int_0^x 1 dt = x$$

3) if  $1 < x < 2$   $x-1 > 0$   $x > 1$

$$\int_{x-1}^x f(t) dt = \int_{x-1}^1 \cancel{f(t) dt} + \int_1^x \cancel{f(t) dt} = \int_{x-1}^1 1 dt = 1 - (x-1) = 2 - x$$

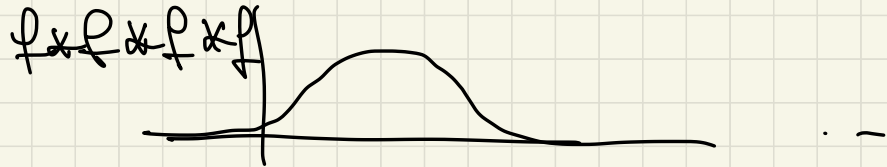
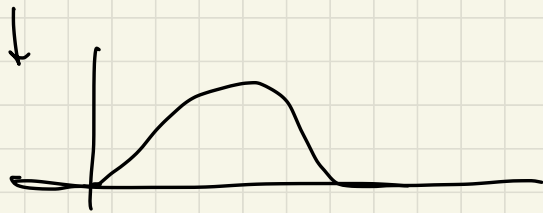
4)  $x > 2$   $x-1 > 1$   $(x-1, x) \subseteq (1, +\infty) \rightarrow \int_{x-1}^x f(t) dt = 0$

$$f * f(x) = \begin{cases} 0 & x < 0 \\ x & 0 \leq x < 1 \\ 2-x & 1 \leq x < 2 \\ 0 & x \geq 2 \end{cases}$$



$f * f$  is CONTINUOUS!

$f * f * f \dots$  becomes more and more regular



Prop If  $X, Y$  are absolutely continuous independent random variables, with density  $f$  and  $g$  then  $X+Y$  is an absolutely cont. random variable with density  $f * g$

Proof  
 $f$  is the density of  $X$  (density = derivative of the cumulative dist. f.)

$$F_X(a) = \int_{-\infty}^a f(x) dx = \mathbb{P}(\omega \in \Omega \mid X(\omega) \leq a)$$

$$\mathbb{P}(\omega \mid Y(\omega) \leq b) = \int_{-\infty}^b g(x) dx = F_Y(b)$$



I want to compute the density of  $X+Y$

$$t \in \mathbb{R} \quad \mathbb{P}(\omega \mid (X+Y)(\omega) \leq t) =$$

$$= \mathbb{P}(\omega \mid \underbrace{X(\omega) \leq x}, \underbrace{Y(\omega) \leq y} \quad x+y \leq t)$$

$$= \int_{\{(x,y) \in \mathbb{R}^2, x+y \leq t\}} f(x) g(y) dx dy = \int_{\mathbb{R}} \int_{-\infty}^{t-x} f(x) g(y) dy dx$$

$$x \in \mathbb{R} \quad y \leq t-x \rightarrow z = y+x \leq t$$

$$= \int_{\mathbb{R}} \int_{-\infty}^t f(x) g(z-x) dz dx = \int_{-\infty}^t \underbrace{\int_{\mathbb{R}} f(x) g(z-x) dx}_{\text{density of } X+Y}$$

$$= \int_{-\infty}^t (f+g)(z) dz$$

$$\boxed{\begin{array}{l} z = x+y \\ x = x \end{array}}$$

(if  $X, Y$  are discrete random variables  
and independent

$$\mathbb{P}((X+Y)(\omega)=t) = \sum_{k \in \mathbb{Z}} \mathbb{P}(X(\omega)=k) \mathbb{P}(Y(\omega)=t-k)$$

# FOURIER TRANSFORM

$$f: \mathbb{R} \rightarrow \mathbb{C} \quad f(x) = f_1(x) + i f_2(x) \quad i = \text{imaginary unit}$$

$$e^{ix} = \cos x + i \sin x \quad \text{for } x \in \mathbb{R} \quad |e^{ix}| = 1$$

$$z \in \mathbb{C} \iff z = x + iy \quad |z| = \sqrt{x^2 + y^2}$$

Take  $f \in L^1(\mathbb{R})$

$$\begin{aligned} \hat{f}(z) &= \text{Fourier transform of } f := \int_{\mathbb{R}} f(y) e^{ixy} dy = \\ &= \int_{\mathbb{R}} \cos(xy) f(y) dy + i \int_{\mathbb{R}} \sin(xy) f(y) dy \end{aligned}$$

$$\hat{f}: \mathbb{R} \rightarrow \mathbb{C}$$

$$\hat{f} \text{ is bounded: } \forall x \in \mathbb{R} \quad |\hat{f}(x)| = \left| \int_{\mathbb{R}} f(y) e^{ixy} dy \right| \leq$$

$$\leq \int_{\mathbb{R}} |f(y)| |e^{ixy}| dy = \int_{\mathbb{R}} |f(y)| dy = \|f\|_1$$

2)  $\hat{f}$  is continuous:  $\hat{f}(x+h) =$

$$= \int_{\mathbb{R}} \underline{f(y)} \underline{\cos((x+h)y)} dy + i \int_{\mathbb{R}} \underline{f(y)} \underline{\sin((x+h)y)} dy$$

$$\xrightarrow{h \rightarrow 0} \int_{\mathbb{R}} f(y) \cos(xy) dy + i \int_{\mathbb{R}} f(y) \sin(xy) dy = \hat{f}(x)$$

$f \in L^1(\mathbb{R}) \longrightarrow \hat{f} : \mathbb{R} \rightarrow \mathbb{C}$  is a continuous bounded function

$$\sup_{x \in \mathbb{R}} |\hat{f}(x)| \leq \|f\|_1$$

Other important properties of FOURIER TRANSFORM:

$$(1) f, g \in L^1(\mathbb{R}) \quad (\widehat{f * g})(x) = \widehat{f}(x) \cdot \widehat{g}(x)$$

(Fourier transform of a convolution of 2 functions is the product of the Fourier transform).

proof:

$$e^{ixy} = e^{ix(y-z)} \cdot e^{ixz}$$

$$(\widehat{f * g})(x) = \int_{\mathbb{R}} (f * g)(y) \cdot e^{ixy} dy =$$

$$= \int_{\mathbb{R}} \int_{\mathbb{R}} f(y-z) g(z) dz e^{ix(y-z)} e^{ixz} dy =$$

$$= \int_{\mathbb{R}} \int_{\mathbb{R}} f(y-z) e^{ix(y-z)} g(z) e^{ixz} dz dy =$$

$z \rightarrow z$   
 $y \rightarrow w = y - z$

$$\begin{aligned}
 &= \int_{\mathbb{R}} \int_{\mathbb{R}} \underbrace{f(w) e^{ixw}} \underbrace{g(z) e^{izx}} dz dw = \\
 &= \underbrace{\int_{\mathbb{R}} f(w) e^{ixw} dw}_{\mathbb{R}} \underbrace{\int_{\mathbb{R}} g(z) e^{izx} dz}_{\mathbb{R}} = \hat{f}(x) \hat{g}(x)
 \end{aligned}$$

(2) Let  $f \in L^1(\mathbb{R})$  such that  $x \cdot f(x) \in L^1(\mathbb{R})$

$$\Rightarrow (\hat{f})'(x) = \frac{d}{dx} \hat{f}(x) = \lim_{h \rightarrow 0} \frac{\hat{f}(x+h) - \hat{f}(x)}{h} =$$

$$= \left[ \underline{\underline{iyf(y)}} \right]^{\wedge}(x)$$

$$\begin{aligned}
 \text{proof} \\
 \frac{d}{dx} \hat{f}(x) &= \frac{d}{dx} \int_{\mathbb{R}} f(y) e^{ixy} dy = \int_{\mathbb{R}} f(y) \frac{d}{dx} e^{ixy} dy =
 \end{aligned}$$

$$= \int_{\mathbb{R}} f(y) i y e^{i x y} dy = \int_{\mathbb{R}} \underbrace{i y f(y)} \cdot e^{i x y} dy$$

$$= \left( \underbrace{i y f(y)} \right)^{\wedge} (x)$$

$$2) \frac{d^2}{dx^2} \hat{f}(x) = \int_{\mathbb{R}} f(y) \frac{d^2}{dx^2} e^{i x y} dy = \int_{\mathbb{R}} f(y) \cdot (i y) (i y) e^{i x y} dy$$

$$= \int_{\mathbb{R}} (-y^2) f(y) e^{i x y} dy = \left( -y^2 f(y) \right)^{\wedge} (x)$$

$$\frac{d^k}{dx^k} \hat{f}(x) = \left( (i y)^k f(y) \right)^{\wedge} (x)$$

3) if  $f$  is differentiable and

$$\lim_{|x| \rightarrow +\infty} f(x) = 0$$

$$\left(\frac{d}{dy} f\right)^{\wedge}(x) = (-ix) \hat{f}(x)$$

proof

$$\left(\frac{d}{dy} f\right)^{\wedge}(x) = \int_{\mathbb{R}} \underbrace{\frac{d}{dy} f(y)} \cdot \underbrace{e^{ixy}} dy =$$

= integration by part formula =  ~~$\left[ f(y) e^{ixy} \right]_{-\infty}^{+\infty} +$~~

$$- \int_{\mathbb{R}} f(y) \cdot \frac{d}{dy} e^{ixy} dy = -ix \int_{\mathbb{R}} f(y) e^{ixy} dy$$



(3')  $f$  is differentiable  $k$ -times and

Order  $|x| \rightarrow \infty$   $f^{(n)}(x) = 0 \quad \forall n \leq k-1$

$$f^{(n)}(x) = \frac{d^n}{dx^n} f(x)$$

$$\left( \frac{d^k}{dy^k} f(y) \right) \Big|_{y=x} = (-ix)^k \cdot \hat{f}(x)$$

$\mathcal{E}_x$ 

$$f(x) = e^{-x^2}$$

$$\hat{f}(x) = \int_{\mathbb{R}} e^{-y^2} e^{ixy} dy = \int_{\mathbb{R}} e^{-y^2} \cos(xy) dy + i \int_{\mathbb{R}} e^{-y^2} \sin(xy) dy$$

How to compute explicitly  $\hat{f}$ ?

$$\frac{d}{dx} (\hat{f}(x)) = \int_{\mathbb{R}} iy e^{-y^2} e^{ixy} dy =$$

$$\left( \frac{d}{dx} \hat{f}(x) = (iy \hat{f}(x))' \right)$$

$$= -\frac{i}{2} \int_{\mathbb{R}} (-2y) e^{-y^2} e^{ixy} dy = \frac{-i}{2} \int_{\mathbb{R}} \frac{d}{dy} (e^{-y^2}) e^{ixy} dy$$

$$= \frac{-i}{2} \left[ \left[ e^{-y^2} e^{ixy} \right]_{-a}^{+a} - \int_{\mathbb{R}} e^{-y^2} \frac{d}{dy} e^{ixy} dy \right] =$$

$$= -\frac{i}{2} \left[ -ix \int_{\mathbb{R}} e^{-y^2} e^{ixy} dy \right] =$$

$$= -\frac{x}{2} \underbrace{\int_{\mathbb{R}} e^{-y^2} e^{ixy} dy}_{\hat{f}(x)}$$

$$\frac{d}{dx} \hat{f}(x) = -\frac{x}{2} \hat{f}(x)$$

$$\frac{1}{\hat{f}(x)} \cdot \frac{d}{dx} \hat{f}(x) = \frac{-x}{2}$$

$$\frac{d}{dx} \log(\hat{f}(x)) = \frac{d}{dx} \left( \frac{-x^2}{4} \right) \implies \square$$

$$f(y) = e^{-y^2}$$

$$\hat{f}(x) = \int_{\mathbb{R}} e^{-y^2} e^{ixy} dy > 0$$

$$\log(\hat{f}(x)) = -\frac{x^2}{4} + c$$

$$e^{\log(\hat{f}(x))} = \hat{f}(x) = e^{-\frac{x^2}{4} + c} = e^{-\frac{x^2}{4}} e^c$$

$$\hat{f}(x) = e^{-\frac{x^2}{4}} e^c$$

$$\hat{f}(0) = e^c = \int_{\mathbb{R}} e^{-y^2} e^{i0 \cdot y} dy = \int_{\mathbb{R}} e^{-y^2} dy = \sqrt{\pi}$$

$$\hat{f}(x) = \sqrt{\pi} e^{-x^2/4}$$

$$f(x) = e^{-x^2}$$